

RECURSIVE FILTERING (i)

Objective:

Given $\left[x(k) = s(k) + n(k) \quad , \quad k=0, 1, \dots, n \right]$

where: $s(n)$ is a random signal with correlation function $R_s(l)$
 $n(n)$ is a white noise process with variance σ_n^2 , zero mean
 $s(n)$ and $n(n)$ are uncorrelated.

Find the best linear-mean-square estimate of $s(n)$, $\hat{s}(n)$
that is find $c_0, \dots, c_n$ so that

$$\hat{s}(n) = \sum_{i=0}^n c_i x(i) \quad \text{by minimizing} \quad E\{|s(n) - \hat{s}(n)|^2\}$$

Obtain $\hat{s}(n)$ recursively:

$$\left[\hat{s}(n) = B_n \hat{s}(n-1) + K_n x(n) \right]$$

Determine conditions for recursive estimation
(Determine values for B_n and K_n .)



RECURSIVE FILTERING (2)

Let $e(n) = s(n) - \hat{s}(n)$, $\sigma_{e(n)}^2 = \text{Var}[e(n)]$

From problem assumptions we have :

$$\underline{R_{xs}(l)} = E\{s(n)x^*(n-l)\} = E\{s(n) \cdot (\hat{s}^*(n-l) + \eta^*(n-l))\} = E\{s(n)s^*(n-l)\} = \underline{R_s(l)} \quad \sigma_{\eta}^2 \delta(l)$$

$$\underline{R_x(l)} = E\{x(n)x^*(n-l)\} = E\{(s(n)+n(n))(\hat{s}^*(n-l)+\eta^*(n-l))\} = E\{s(n)s^*(n-l)\} + E\{n(n)\eta^*(n-l)\} = \underline{R_s(l)} + \underline{R_{\eta}(l)}$$

$$\underline{R_{nx}(0)} = E\{x(n)\eta^*(n)\} = E\{(s(n)+n(n)) \cdot \eta^*(n)\} = E\{n(n)\eta^*(n)\} = \underline{\sigma_{\eta}^2} = \underline{R_{\eta}(0)}$$

Apply orthogonality condition: $E\{e(n)x^*(n-l)\} = 0$, $l=0,1,\dots,n$

$l=0$: $E\{e(n) [\hat{s}^*(n) + \eta^*(n)]\} = \underbrace{E\{e(n)\hat{s}^*(n)\}}_{\sigma_{e(n)}^2} + E\{e(n)\eta^*(n)\} = \sigma_{e(n)}^2 + E\{(s(n) - \hat{s}(n))\eta^*(n)\}$

$$= \sigma_{e(n)}^2 + E\{[s(n) - B_n \hat{s}(n-1) - K_n x(n)] \cdot \eta^*(n)\} = 0 \Rightarrow \sigma_{e(n)}^2 - \overset{K_n}{E\{x(n)\eta^*(n)\}} = 0$$

$$\Rightarrow \sigma_{e(n)}^2 - K_n R_{\eta}(0) = 0 \Rightarrow \left[K_n = \frac{\sigma_{e(n)}^2}{\sigma_{\eta}^2} \right] \text{ real and positive.}$$



RECURSIVE FILTERING (3)

$l > 0$:

$$E\{e(n)x^*(n-l)\} = E\{s(n)x^*(n-l)\} - B_n E\{\hat{s}(n-1)x^*(n-l)\} - K_n E\{x(n)x^*(n-l)\} = 0$$

$$\Rightarrow R_s(l) - B_n E\{[s(n-1) - e(n-1)]x^*(n-l)\} - K_n (R_s(l) + R_n(l)) = 0 \Rightarrow$$

$\leftarrow 0 \ (l > 0)$

$$\Rightarrow (1 - K_n) \cdot R_s(l) - B_n R_s(l-1) - \overset{\leftarrow 0}{0} = 0 \Rightarrow$$

$$\Rightarrow \left[R_s(l) - \frac{B_n}{1 - K_n} R_s(l-1) = 0 \right], \quad l > 0$$

[NOTE: since this is a recursive filter the condition $E\{e(n-1)x^*(n-l)\} = 0$ has been "achieved" in the previous iteration]

Solution of the above difference equation provides

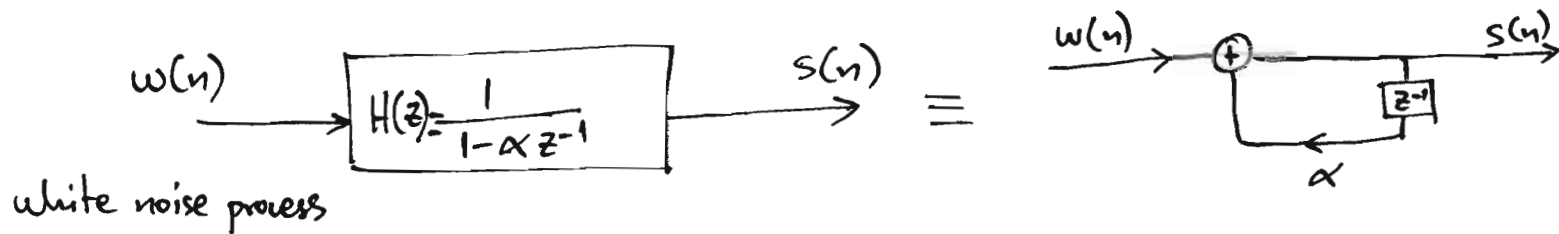
$$\left[R_s(l) = R_s(0) \alpha^l, \quad l > 0 \right] \quad \text{where} \quad \left[\alpha = \frac{B_n}{1 - K_n} \right]$$

exponential form of correlation function.



RECURSIVE FILTERING (4)

Consider the signal $s(n) = \alpha s(n-1) + w(n)$



i.e. first order autoregressive model. Then for this signal it is easy to show that

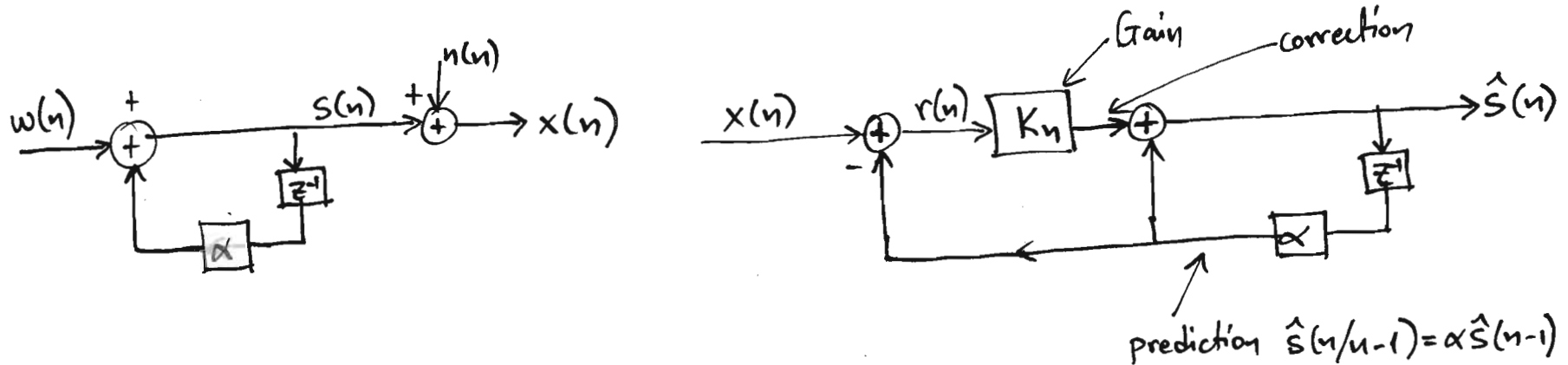
$$R_s(l) = \sigma_w^2 \cdot \frac{\alpha^l}{1 - |\alpha|^2} = R_s(0) \cdot \alpha^l$$

Therefore, for a signal $s(n)$ generated as above a recursive solution to the optimal filtering as defined previously is possible.

RECURSIVE FILTERING (5)

So, $\hat{s}(n) = B_n \hat{s}(n-1) + K_n x(n) = \alpha(1-K_n) \hat{s}(n-1) + K_n x(n)$

or $\left[\underbrace{\hat{s}(n) = \alpha \hat{s}(n-1)}_{\text{forward predictor}} + \underbrace{K_n [x(n) - \alpha \hat{s}(n-1)]}_{\text{Correction term}} \right], K_n = \frac{\sigma_e^2}{\sigma_n^2}$



$r(n)$: residual error $\left\{ \begin{array}{l} \text{white non-stationary, represents the new information in observations} \\ \text{white stationary at steady state } (n \rightarrow \infty); \text{ innovations process of } x(n) \end{array} \right.$

RECURSIVE FILTERING (6)

To define K_n better let us derive an expression for $\sigma_{e(n)}^2$

From orthogonality principle: $\sigma_{e(n)}^2 = E\{s(n) \cdot e^*(n)\} = \dots = (1 - K_n)[\sigma_s^2 - |\alpha|^2 E\{s(n-1)\hat{s}^*(n-1)\}]$

for $n=0$ (assuming $\hat{s}(-1)=0$) : $\sigma_{e(0)}^2 = \frac{\sigma_s^2}{1 + \sigma_s^2/\sigma_w^2} = \frac{\sigma_w^2}{1 + \sigma_w^2/\sigma_n^2}$

for $n > 0$ it can be shown that $\left[\sigma_{e(n)}^2 = \frac{\sigma_w^2 + |\alpha|^2 \sigma_{e(n-1)}^2}{\sigma_n^2 + \sigma_w^2 + |\alpha|^2 \sigma_{e(n-1)}^2} \cdot \sigma_n^2 \right]$: Recursive expression.

So $\left[K_n = \frac{\sigma_w^2 + |\alpha|^2 \sigma_{e(n-1)}^2}{\sigma_n^2 + \sigma_w^2 + |\alpha|^2 \sigma_{e(n-1)}^2} \right]$, if $\sigma_w^2 \gg \sigma_n^2 \Rightarrow K_n \rightarrow 1 \Rightarrow \underline{\hat{s}(n) = x(n)}$

$K_n, \sigma_{e(n)}^2$ are computed recursively but can be computed a-priori since they do not depend on $x(n)$

$K_0 = \frac{\sigma_w^2}{\sigma_n^2 + \sigma_w^2}$



RECURSIVE FILTERING (7)

Example: Constant signal in noise

$$\text{let } x(n) = s + n(n) \quad n=0, 1, \dots, N$$

where, s is a constant ($w(n)=0$, $\alpha=1$)

Then, $s(n) = s(n-1) = s$ and.

$$\hat{s}(n) = \hat{s}(n-1) + K_n (x(n) - \hat{s}(n-1))$$

$$\text{Since } \begin{matrix} \sigma_w^2 = 0 \\ \alpha = 1 \end{matrix}, \quad \sigma_{e(n)}^2 = \frac{\sigma_{e(n-1)}^2}{\sigma_n^2 + \sigma_{e(n-1)}^2} \cdot \sigma_n^2 = \frac{\sigma_{e(n-1)}^2}{1 + \frac{\sigma_{e(n-1)}^2}{\sigma_n^2}} < \sigma_{e(n-1)}^2 < \sigma_{e(n-2)}^2$$

Thus, as $n \rightarrow \infty$ $\sigma_{e(n)}^2 \rightarrow 0$ and $\hat{s}(n) \rightarrow s$

RECURSIVE PREDICTION (1)

Given $[X(k) = s(k) + n(k), \quad k=0, 1, \dots, n]$

estimate, (predict) $s(n+1)$ in the minimum mean square sense.
using a linear estimator

where we assume that $s(n) = \alpha s(n-1) + w(n)$,
 $s(n), n(n)$ are uncorrelated etc....

Define the estimate
let the estimate

$$\hat{s}(n+1/n) = \alpha \hat{s}(n)$$

← this is the optimal estimator.

and the error

$$e(n+1/n) = s(n+1) - \hat{s}(n+1/n)$$

Then it can be shown that $E\{x(n-l)e^*(n+1/n)\} = 0 \quad l=0, 1, 2, \dots, n$

and $\sigma_{e(n+1/n)}^2 = E\{s(n+1)e^*(n+1/n)\} = |\alpha|^2 \sigma_{e(n)}^2 + \sigma_w^2$



RECURSIVE PREDICTION (2)

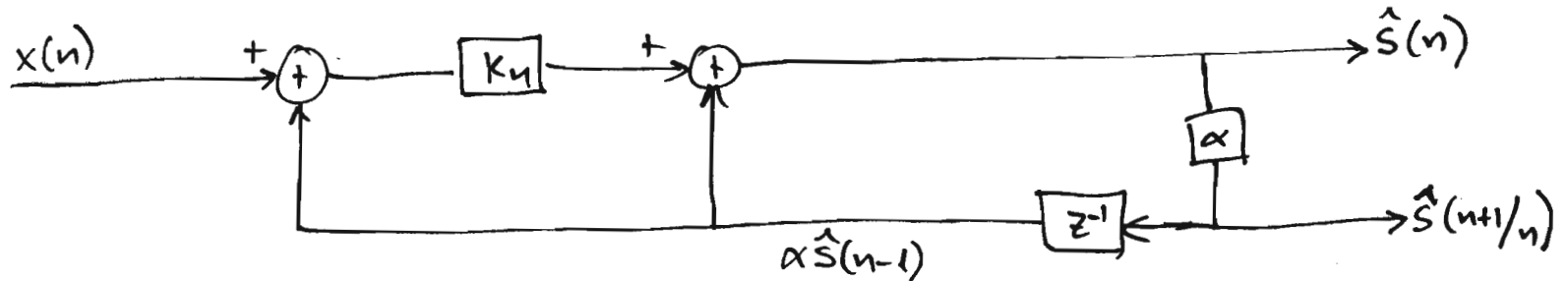
From previous results $\left[\hat{s}(n) = \alpha \hat{s}(n-1) + K_n [x(n) - \alpha \hat{s}(n-1)] \right]$, $\hat{s}(n+1/n) = \alpha \hat{s}(n)$

So $\hat{s}(n+1/n) = \alpha \left(\alpha \hat{s}(n-1) + K_n [x(n) - \alpha \hat{s}(n-1)] \right) \Rightarrow$

$\Rightarrow \left[\hat{s}(n+1/n) = \alpha \cdot \hat{s}(n/n-1) + \alpha K_n [x(n) - \hat{s}(n/n-1)] \right]$

$K_n = \frac{\sigma_e^2(n) - \sigma_w^2}{1 + \sigma_e^2(n)} = \frac{\sigma_e^2(n)}{\sigma_e^2(n) + \sigma_w^2}$
i.e., same as before.

Realization
1

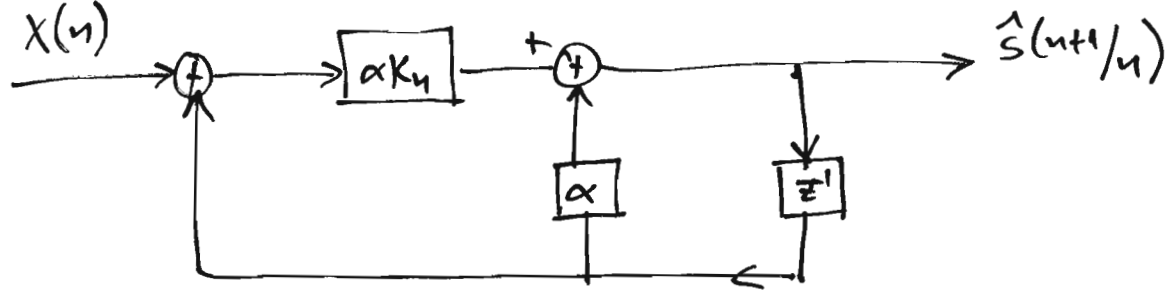


System to perform simultaneous filtering and prediction.

RECURSIVE PREDICTION (3)

① Direct realization of predictor. (recursive realization)

Realization
②



$$\alpha K_n = \frac{\sigma_{e(n+1/n)}^2 - \sigma_w^2}{\alpha^* \sigma_n^2}, \quad \sigma_{e(n+1/n)}^2 = \sigma_w^2 + \frac{|\alpha|^2 \sigma_n^2 \sigma_{e(n/n-1)}^2}{\sigma_n^2 + \sigma_{e(n/n-1)}^2}$$

* Note: for $\sigma_n^2 \rightarrow 0$ it can be shown that $K_n \rightarrow 1$ and $\hat{s}(n+1/n) = \alpha x(n)$

RECURSIVE FILTERING

- ③ The generalization of the presented recursive procedure to vector valued random processes is known as the Kalman filter
- ③ The recursive filter is a special case of the Wiener filter that is causal but not time invariant. It begins by processing a single observation and continues to use more samples of the observation process to estimate the desired signal as those samples become available.
- ③ The filter parameters are also computed recursively but are independent of the observed data so they can be computed "off line."
- ③ Asymptotically (or after a large number of recursions) the filter approaches steady state and becomes equivalent to the Infinite length Wiener filter.....


Generalization of Kalman filter.

Let
$$s(n) = \sum_{k=1}^p a(k) s(n-k) + w(n)$$

$$y(n) = s(n) + v(n)$$

Let
$$\underline{s}(n) = \begin{bmatrix} s(n) \\ s(n-1) \\ \vdots \\ s(n-p+1) \end{bmatrix} \quad p \times 1$$

Then
$$\underline{s}(n) = \begin{bmatrix} a(1) & a(2) & \dots & a(p) \\ 1 & 0 & \dots & 0 \\ 0 & 1 & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & \vdots & \dots & 1 \end{bmatrix} \underline{s}(n-1) + \begin{bmatrix} 1 \\ 0 \\ \vdots \\ 0 \end{bmatrix} w(n)$$

$$y(n) = [1 \ 0 \ 0 \ \dots \ 0] \underline{s}(n) + v(n)$$

or state eqn $\underline{s}(n) = \underline{A} \underline{s}(n-1) + \underline{w}(n)$ → state transition matrix

measurement (observable) $y(n) = \underline{C} \underline{s}(n) + v(n)$ ← measurement noise) mean let

$$E\{\underline{w}(n)w^H(k)\} = Q_w(n) \delta(n-k) \quad E\{v(n)v^H(k)\} = Q_v(n) \delta(n-k)$$

Optimal estimate
Given measurements up to time n

$$\hat{\underline{s}}(n) = \underline{A} \hat{\underline{s}}(n-1) + \underline{K} [y(n) - \underline{C} \underline{A} \hat{\underline{s}}(n-1)]$$

Non-stochastic case

$$\underline{s}(n) = \sum_{k=1}^p a(k) \underline{s}(n-k)$$

$$\underline{s}(n) = \underline{A}(n-1) \underline{s}(n-1) + \underline{w}(n)$$

$$y(n) = \underline{C}(n) \underline{s}(n) + v(n)$$

$$\hat{\underline{s}}(n) = \underline{A}(n-1) \hat{\underline{s}}(n-1) + \underline{K}(n) [y(n) - \underline{C}(n) \underline{A}(n-1) \hat{\underline{s}}(n-1)]$$

DISCRETE KALMAN FILTER

• $x(k) = s(k) + \eta(k)$

Known statistics

$x(k), s(k)$ jointly stationary
 $s(k), \eta(k)$ independent or
at least uncorrelated

- Optimum FIR Wiener Filter $\rightarrow$ Normal equations (Wiener-Hopf equations)
- Optimum IIR Wiener Filter $\rightarrow$ Non-causal solution non-realizable but easy to design $\rightarrow$ $\left\{ \begin{array}{l} \text{performance} \\ \text{bounds} \end{array} \right.$
- Causal ^{IIR} Wiener filtering $\rightarrow$ Spectral factorization $\rightarrow$ realizable but more difficult to design.
- Recursive filtering $\rightarrow$ discrete Kalman Filter
Applicable to non-stationary situations as well

AR(1) modeling of $s(n)$: $s(n) = a s(n-1) + w(n)$

• ~~AR(1)~~ Recursive solution - stationary case

$$\hat{s}(n) = a \hat{s}(n-1) + K_n [x(n) - a \hat{s}(n-1)]$$

Recursive solution - non-stationary case

$$\hat{s}(n) = a_{n-1} \hat{s}(n-1) + K_n [x(n) - a_{n-1} \hat{s}(n-1)]$$

Define

$$\begin{aligned} e(n/n) &= s(n) - \hat{s}(n/n) && \text{filtering error} \\ e(n/n-1) &= s(n) - \hat{s}(n/n-1) && \text{prediction error} \end{aligned}$$

$P(n/n)$ filtering error covariance matrix

$P(n/n-1)$ prediction error covariance matrix

The Discrete Kalman Filter

State equation $\underline{s}(n) = \underline{A}(n-1) \underline{s}(n-1) + \underline{w}(n)$

Observation equation $\underline{y}(n) = \underline{C}(n) \underline{s}(n) + \underline{v}(n)$

Initialization:

$$\begin{aligned} \hat{\underline{s}}(0/0) &= E\{\underline{s}(0)\} \\ \underline{P}(0/0) &= E\{\underline{s}(0) \underline{s}^H(0)\} \end{aligned}$$

Computation : for $n=1, 2, \dots$

$$\hat{\underline{s}}(n/n-1) = \underline{A}(n-1) \hat{\underline{s}}(n-1/n-1)$$

$$\underline{P}(n/n-1) = \underline{A}(n-1) \underline{P}(n-1/n-1) \underline{A}^H(n-1) + \underline{Q}_w(n)$$

$$\underline{K}(n) = \underline{P}(n/n-1) \underline{C}^H(n) [\underline{C}(n) \underline{P}(n/n-1) \underline{C}^H(n) + \underline{Q}_v(n)]^{-1}$$

$$\hat{\underline{s}}(n/n) = \hat{\underline{s}}(n/n-1) + \underline{K}(n) [\underline{y}(n) - \underline{C}(n) \hat{\underline{s}}(n/n-1)]$$

$$\underline{P}(n/n) = [\underline{I} - \underline{K}(n) \underline{C}(n)] \underline{P}(n/n-1)$$