

# Handoff Rate Analysis in Heterogeneous Cellular Networks: a Stochastic Geometric Approach

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## ABSTRACT

Horizontal and vertical handoffs are important ramifications of user mobility in multi-tier heterogeneous cellular networks. They directly affect the signaling overhead and quality of calls in the system. However, they are difficult to analyze due to the irregularly shaped network topologies introduced by multiple tiers of cells. In this work, a stochastic geometric analysis framework on user mobility is proposed, to capture the spatial randomness and various scales of cell sizes in different tiers. We derive theoretical expressions for the rates of all handoff types experienced by an active user with arbitrary movement trajectory. Empirical study using real user mobility trace data and extensive simulation are conducted, demonstrating the correctness and usefulness of our analysis.

## Categories and Subject Descriptors

C.2.1 [Network Architecture and Design]: Wireless communication

## Keywords

Cellular networks; mobility; handoff; stochastic geometry; analytic geometry

## 1. INTRODUCTION

Traditional single-tier macro-cellular networks provide wide coverage for mobile user equipments (UEs), but they are insufficient to satisfy the exploding demand for high bandwidth access driven by modern mobile traffic, such as multimedia transmissions and cloud computing tasks. One effective means to increase network capacity is to provide more serving stations within a geographical area, i.e., installing a diverse set of small-cells such as femtocells [14] and WiFi hotspots [12], overlaying the macrocells, to form a multi-tier heterogeneous cellular network. Each small-cell is equipped with a shorter-range and lower-cost base station (BS) or access point (AP), to provide nearby UEs with

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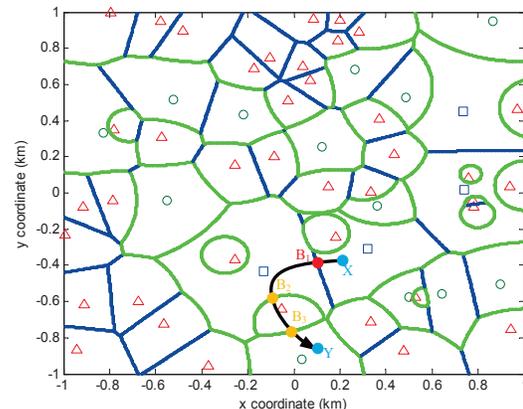


Figure 1: An example of a three-tier cellular network. Tier-1, 2, and 3 BSs are represented by squares, circles, and triangles respectively; blue curves show intra-tier cell boundaries; green curves show inter-tier cell boundaries. A UE starts a call at  $X$  and terminates it at  $Y$ . It experiences one horizontal handoff at  $B_1$  and two vertical handoffs at  $B_2$  and  $B_3$ .

higher-bandwidth network access with lower power usage, and to offload data traffic from macrocells. The commercial deployment of small-cells has attracted increasing attention in recent years. For example, AT&T Inc. now supplies a femtocell product [1], and it has also deployed WiFi APs in a number of metropolitan areas with dense population [2].

In the presence of multiple tiers of cells, however, mobile UEs may experience internetworking issues among different tiers. In particular, vertical handoffs (i.e., handoffs made between two BSs in different tiers) are introduced [41]. Compared with horizontal handoffs (i.e., handoffs made between two BSs in the same tier), vertical handoffs have a more complicated impact on both the UEs and the overall system. Additional risks are present during channel setup and tear down when a vertical handoff is made, such as (1) extra traffic latency; (2) additional network signaling; (3) more UE power consumption due to simultaneously active network interface to multiple tiers; and (4) higher risk in call drops or degraded quality of service (QoS) caused by the lack of radio resource after handoffs. Furthermore, vertical handoffs may be classified into inter-RAT (radio access technology) handoffs (e.g., handoffs made between LTE access and WiFi access) and intra-RAT handoffs, where the former could cause worse performance degradation on UEs [3].

The handoff rate is defined as the expected number of handoffs experienced by one UE per unit time, which directly affects the signaling overhead in the system and QoS of UEs. As a prerequisite to performance evaluation and system design in heterogeneous cellular networks, it is essential to quantify the rates of different handoff types. However, a study on handoff rates in heterogeneous cellular networks will inevitably be challenged by the irregularly shaped multi-tier network topologies introduced by the small-cell structure. An example topology with three tiers of BSs is shown in Fig. 1. First, BSs are spread irregularly, sometimes in an anywhere plug-and-play manner, leading to a high level of spatial randomness. Second, different tiers of cells are equipped with BSs communicating at different power levels, causing various scales of cell sizes. As a consequence, it is difficult to characterize the cell boundaries and to track boundary crossings made by UEs (i.e., handoffs) in the system. Few previous works have resolved the above challenges.

In this work, we contribute to user mobility modeling by providing new technical tools to quantify the rates of horizontal and vertical handoffs, under random multi-tier BSs, arbitrary user movement trajectory, and flexible user-BS association. A new stochastic geometric analysis framework on user mobility is proposed. In this framework, different tiers of BSs are modeled as Poisson point processes (PPPs) to capture their spatial randomness. To model flexible scaling of cell sizes in different tiers, we consider the biased user association scheme [10, 27, 32, 40], in which each tier of BSs is assigned an association bias value, and a UE is associated with a BS that provides the largest biased received power. Through stochastic and analytic geometric analysis, we derive exact expressions for the rates of all handoff types experienced by an active UE with arbitrary movement trajectory.

We confirm our theoretical analysis through an empirical study using the Yonsei Trace [17]. The trace provides a large data set, accumulating fine-grained mobility data from commercial mobile phones in an 8-month period. Numerical studies using the empirical trace data set, together with further simulation, demonstrate the correctness and usefulness of our analytical conclusions.

The rest of the paper is organized as follows. In Section 2, we discuss the relation between our work and prior works. In Section 3, we describe the system model. In Section 4, we present our contributions in handoff rates derivations. In Section 5, we present empirical study with the Yonsei Trace as well as simulation. Finally, conclusions are given in Section 6.

## 2. RELATED WORKS

### 2.1 Mobility Modeling Based on Queueing Systems

One common category of previous works employ queueing systems to model heterogeneous cellular networks. In this case, cells are modeled as queues, active users are modeled as units in the queues, and handoffs correspond to unit transfers among queues. Ghosh *et al.* [22] studied the single-cell scenario using an  $M/G/\infty$  queue. Kirsal *et al.* [28] studied one WLAN cell overlaying one 3G cell, and a two-queue model is proposed accordingly. For multicell scenarios, queueing network models have been employed in [6, 9, 15, 19, 35]. However, none of these works explicitly

modeled the geometric patterns of cell shapes in heterogeneous networks.

### 2.2 Geometric Pattern Study

In order to characterize the geometric patterns of cellular network topologies, a second category of works model the shape of cells, mostly in non-random regular grids. Zonoozi and Dassanayake [45] modeled a one-tier cellular network as a hexagonal grid. Anpalagan and Katzela [4] studied a two-tier cellular network by modeling small-cells as hexagons, and each macrocell as a cluster of neighbouring small-cells. Shenoy and Hartpence [38] studied a two-tier network by modeling WLAN small-cells as squares, and macrocells as larger squares, each covering  $5 \times 5$  WLAN cells. Hasib and Fapojuwo [24] studied a two-tier cellular network including one hexagonal macrocell and a predetermined  $N$  circular microcells. Lin *et al.* [31] conducted a pioneering study on the user mobility in one-tier macro cellular network considering randomly distributed BSs. Macrocells were modeled as a standard Poisson Voronoi. However, in [31], the authors did not consider multi-tier BSs with different scales of cell sizes. To the best of our knowledge, ours is the first work studying user mobility in *multi-tier* heterogeneous cellular networks that captures their *random* geometric patterns.

### 2.3 Real-world Trace Study

Another important category of related works employ empirical traces to investigate user mobility. Kotz *et al.* [25, 29] studied user mobility patterns on the Dartmouth campus. McNett and Voelker [34] characterized the mobility and access patterns of hand-held PDA users on the UCSD campus, and a campus waypoint model was proposed to characterize the trace. Halepovic and Williamson [23] studied mobility parameters such as the number of calls initiated per user, call inter-arrival time, and the number of cell sites visited per user, based on data traffic traces of a regional CDMA2000 cellular network. Rhee *et al.* [37] concluded that human walk patterns contain statistically similar features observed in Levy walks, based on a large daily GPS trace set accumulated in 5 different places in US and Korea. In [16], empirical study on spatial and temporal mobility patterns of the Yonsei Trace [17] was conducted, in order to predict users' future position precisely. Ficek and Kencl [21] proposed inter-call mobility model to locate users' position between calls based on the trace accumulated in a trip between San Jose and San Francisco. Baumann *et al.* [11] predicted user arrival and residence times in the system through extracting important parameters from the trace accumulated by Nokia Research.

These works based on real-world traces study are practically valuable for system evaluation and design. However, they are insufficient to provide in-depth analytical modeling of handoff and data rates. In our work, we use the Yonsei Trace [16, 17] to demonstrate the correctness and usefulness of our theoretical results.

### 2.4 Handoff and Association Decision Algorithms

Orthogonal to the scope of this work, there is a large body of previous works that study handoff timing algorithms, without considering the random geometric patterns of UEs and BSs. One type of handoff decision algorithms employ a threshold comparison of one or several specific metrics,

(e.g., received signal strength, network loading, bandwidth, and so on) to derive handoff decisions [30, 33, 36, 44]. Another type uses dynamic programming (DP) [42] or artificial intelligence techniques (e.g., fuzzy logic [26]) to improve the effectiveness of handoff procedures.

### 3. SYSTEM MODEL

#### 3.1 Multi-tier Cellular Network

We consider a heterogeneous cellular network with spatially randomly distributed  $K$  tiers of BSs. Let  $\mathcal{K} = \{1, 2, \dots, K\}$ . In order to characterize the random spatial patterns of BSs, we use the conventional assumption that each tier of BSs independently form a homogeneous Poisson point process (PPP) in two-dimensional Euclidean space  $\mathbb{R}^2$  [7, 8, 18, 27, 32, 39, 40]. Let  $\Phi_k$  denote the PPP corresponding to tier- $k$  BSs, and let  $\lambda_k$  be its intensity.

#### 3.2 Biased User Association

Different tiers of BSs transmit at different power levels. Let  $P_k$  be the transmission power of tier- $k$  BSs, which is a given parameter. If  $P_t(\mathbf{x})$ , for  $P_t(\mathbf{x}) \in \{P_1, P_2, \dots, P_K\}$ , is the transmission power from a BS at  $\mathbf{x}$  and  $P_r(\mathbf{y})$  is the received power at  $\mathbf{y}$ , we have  $P_r(\mathbf{y}) = \frac{P_t(\mathbf{x})}{\alpha|\mathbf{x}-\mathbf{y}|^\gamma}$ , where  $\alpha|\mathbf{x}-\mathbf{y}|^\gamma$  is the propagation loss function with  $\gamma > 2$ .

In order to capture various scales of different cell sizes, the biased user association is studied in this work [27, 32, 40]. Given that a UE is located at  $\mathbf{y}$ , it associates itself with the BS that provides the maximum *biased received power* as follows:

$$\mathcal{BS}(\mathbf{y}) = \arg \max_{\mathbf{x} \in \Phi_k, \forall k} B_k P_k |\mathbf{x} - \mathbf{y}|^{-\gamma}, \quad (1)$$

where  $\mathcal{BS}(\mathbf{y})$  denotes the location of the BS chosen for the UE,  $P_k |\mathbf{x} - \mathbf{y}|^{-\gamma}$  is the received power from a tier- $k$  BS located at  $\mathbf{x}$ , and  $B_k$  is the association bias, indicating the received power preference of UEs toward tier- $k$  BSs.  $B_k$  may be different in different tiers, mainly because (1) different radio access technologies may require different received power levels, and (2) some tiers could be assigned larger values of  $B_k$ , in order to offload data traffic from other tiers. As a consequence, the resultant cell splitting forms a generalized Dirichlet tessellation, or weighted Poisson Voronoi [5], an example of which is shown in Fig. 1. Let  $\mathbf{T}^{(1)}$  denote the overall cell boundaries, and let  $\mathbf{T}_{kj}^{(1)}$  denote the boundaries of tier- $k$  cells and tier- $j$  cells, which is also referred to as *type  $k$ - $j$  cell boundaries* in this paper. Note that  $\mathbf{T}_{kj}^{(1)}$  and  $\mathbf{T}_{jk}^{(1)}$  are equivalent (type  $k$ - $j$  cell boundaries and type  $j$ - $k$  cell boundaries are equivalent).

Note that for  $B_1, B_2, \dots, B_K$ , their effects remain the same if we multiply all of them by a same positive constant.

For presentation convenience, we define  $\beta_{kj} = \left(\frac{P_k B_k}{P_j B_j}\right)^{1/\gamma}$ . Clearly,  $\beta_{kj} = \frac{1}{\beta_{jk}}$ .

Let  $A_k$  denote the probability that a UE associates itself with a tier- $k$  BS. As derived in [27], we have

$$A_k = \frac{\lambda_k (P_k B_k)^{\frac{2}{\gamma}}}{\sum_{j=1}^K \lambda_j (P_j B_j)^{\frac{2}{\gamma}}}. \quad (2)$$

#### 3.3 UE Trajectory and Handoff Rate

We aim to study the rates of all types of handoffs of some active UE moving in the network. Let  $\mathcal{T}_0$  denote the trajectory of the UE, which is of finite length. The number of handoffs the UE experiences is equal to the number of intersections of  $\mathcal{T}_0$  and  $\mathbf{T}^{(1)}$ , which is denoted by  $\mathcal{N}(\mathcal{T}_0, \mathbf{T}^{(1)})$ . In this paper, a handoff made from a tier- $k$  cell to a tier- $j$  cell is called a *type  $k$ - $j$  handoff*. The number of type  $k$ - $j$  handoffs is denoted by  $\mathcal{N}_{kj}(\mathcal{T}_0, \mathbf{T}_{kj}^{(1)})$ .

If  $j \neq k$ , a type  $k$ - $j$  (vertical) handoff is **not** equivalent to a type  $j$ - $k$  handoff. When the UE crosses type  $k$ - $j$  boundary, either a type  $k$ - $j$  or a type  $j$ - $k$  handoff is made, depending on the moving direction. Thus, the number of type  $k$ - $j$  plus type  $j$ - $k$  handoffs is equal to the number of intersections of  $\mathcal{T}_0$  and  $\mathbf{T}_{kj}^{(1)}$ , which is denoted by  $\mathcal{N}(\mathcal{T}_0, \mathbf{T}_{kj}^{(1)})$ . In other words, we have  $\mathcal{N}(\mathcal{T}_0, \mathbf{T}_{kj}^{(1)}) = \mathcal{N}_{kj}(\mathcal{T}_0, \mathbf{T}_{kj}^{(1)}) + \mathcal{N}_{jk}(\mathcal{T}_0, \mathbf{T}_{kj}^{(1)})$ .

If  $j = k$ ,  $\mathcal{N}(\mathcal{T}_0, \mathbf{T}_{kk}^{(1)}) = \mathcal{N}_{kk}(\mathcal{T}_0, \mathbf{T}_{kk}^{(1)})$  indicates the number of type  $k$ - $k$  (horizontal) handoffs.

In Section 4, we aim to study all types of handoff rates, which correspond to the expected numbers of handoffs experienced by the active UE per unit time.

### 4. HANDOFF RATE ANALYSIS IN MULTI-TIER CELLULAR NETWORKS

The proposed analysis of handoff rates consists of a progressive sequence of four components, which are described in the following subsections.

#### 4.1 Length Intensity of Cell Boundaries

Handoffs occur at the intersections of the active UE's trajectory with cell boundaries. In order to track the number of intersections, we need to first study the length intensity of cell boundaries  $\mathbf{T}^{(1)}$  (resp.  $\mathbf{T}_{kj}^{(1)}$ ), which is defined as the expected length of  $\mathbf{T}^{(1)}$  (resp.  $\mathbf{T}_{kj}^{(1)}$ ) in a unit square. Higher length intensity of cell boundaries leads to greater opportunities for boundary crossing, and thus higher handoff rates.

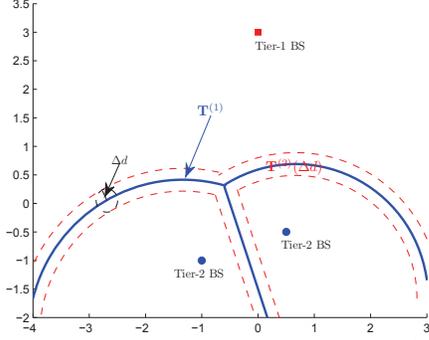
The cell boundaries  $\mathbf{T}^{(1)}$  is a fiber process [43] generated by  $\Phi_1, \Phi_2, \dots, \Phi_K$ .  $\mathbf{T}^{(1)}$  also corresponds to the set of points on  $\mathbb{R}^2$ , where a same biased power level is received from two nearby BSs, and this biased received power level is no less than those from any other BSs. Mathematically, we have

$$\mathbf{T}^{(1)} = \left\{ \mathbf{x} \mid \forall k, j \in \mathcal{K}, \exists \mathbf{x}_1 \in \Phi_k, \mathbf{x}_2 \in \Phi_j, \mathbf{x}_1 \neq \mathbf{x}_2, \right. \\ \left. \text{s.t. } P_r = \frac{P_k B_k}{|\mathbf{x}_1 - \mathbf{x}|^\gamma} = \frac{P_j B_j}{|\mathbf{x}_2 - \mathbf{x}|^\gamma}, \text{ and} \right. \\ \left. \forall i \in \mathcal{K}, \mathbf{y} \in \Phi_i, P_r \geq \frac{P_i B_i}{|\mathbf{y} - \mathbf{x}|^\gamma} \right\}. \quad (3)$$

Similarly,  $\mathbf{T}_{kj}^{(1)}$  can be expressed as

$$\mathbf{T}_{kj}^{(1)} = \left\{ \mathbf{x} \mid \exists \mathbf{x}_1 \in \Phi_k, \mathbf{x}_2 \in \Phi_j, \mathbf{x}_1 \neq \mathbf{x}_2, \right. \\ \left. \text{s.t. } P_r = \frac{P_k B_k}{|\mathbf{x}_1 - \mathbf{x}|^\gamma} = \frac{P_j B_j}{|\mathbf{x}_2 - \mathbf{x}|^\gamma}, \text{ and} \right. \\ \left. \forall i \in \mathcal{K}, \mathbf{y} \in \Phi_i, P_r \geq \frac{P_i B_i}{|\mathbf{y} - \mathbf{x}|^\gamma} \right\}. \quad (4)$$

Note that  $\bigcup_{k=1}^K \bigcup_{j=k}^K \mathbf{T}_{kj}^{(1)} = \mathbf{T}^{(1)}$ .



**Figure 2:** The blue bold curves show  $\mathbf{T}^{(1)}$ ; and the region within red dashed curves shows  $\mathbf{T}^{(2)}(\Delta d)$ .

Let  $\mu_1(\mathbf{T}^{(1)})$  denote the length intensity of  $\mathbf{T}^{(1)}$ , which is the expected length of  $\mathbf{T}^{(1)}$  in a unit square<sup>1</sup> [43]:

$$\mu_1(\mathbf{T}^{(1)}) = \mathbb{E} \left( \left| \mathbf{T}^{(1)} \cap [0, 1]^2 \right|_1 \right), \quad (5)$$

where  $|L|_1$  denotes the length of  $L$  (i.e., one-dimensional Lebesgue measure of  $L$ ). Similarly, let  $\mu_1(\mathbf{T}_{kj}^{(1)})$  denote the length intensity of  $\mathbf{T}_{kj}^{(1)}$ :

$$\mu_1(\mathbf{T}_{kj}^{(1)}) = \mathbb{E} \left( \left| \mathbf{T}_{kj}^{(1)} \cap [0, 1]^2 \right|_1 \right). \quad (6)$$

Note that we have  $\mu_1(\mathbf{T}^{(1)}) = \sum_{k=1}^K \sum_{j=k}^K \mu_1(\mathbf{T}_{kj}^{(1)})$ .

## 4.2 $\Delta d$ -Extended Cell Boundaries

It is difficult to directly quantify the one-dimensional measures  $\mu_1(\mathbf{T}^{(1)})$  and  $\mu_1(\mathbf{T}_{kj}^{(1)})$  on the two-dimensional plane. Instead, we first introduce the  $\Delta d$ -extended cell boundaries, which extends the one-dimensional measures to two-dimensional measures.

The  $\Delta d$ -extended cell boundaries of  $\mathbf{T}^{(1)}$ , denoted by  $\mathbf{T}^{(2)}(\Delta d)$  is defined as

$$\mathbf{T}^{(2)}(\Delta d) = \left\{ \mathbf{x} \mid \exists \mathbf{y} \in \mathbf{T}^{(1)}, \text{ s.t. } |\mathbf{x} - \mathbf{y}| < \Delta d \right\}. \quad (7)$$

In other words,  $\mathbf{T}^{(2)}(\Delta d)$  is the  $\Delta d$ -neighbourhood of  $\mathbf{T}^{(1)}$ . A point is in  $\mathbf{T}^{(2)}(\Delta d)$  iff its (shortest) distance to  $\mathbf{T}^{(1)}$  is less than  $\Delta d$ , as shown in Fig. 2. Similarly, we define  $\mathbf{T}_{kj}^{(2)}(\Delta d)$  as the  $\Delta d$ -extended cell boundaries of  $\mathbf{T}_{kj}^{(1)}$  (i.e.,  $\Delta d$ -neighbourhood of  $\mathbf{T}_{kj}^{(1)}$ ):

$$\mathbf{T}_{kj}^{(2)}(\Delta d) = \left\{ \mathbf{x} \mid \exists \mathbf{y} \in \mathbf{T}_{kj}^{(1)}, \text{ s.t. } |\mathbf{x} - \mathbf{y}| < \Delta d \right\}. \quad (8)$$

The area intensity of  $\mathbf{T}^{(2)}(\Delta d)$  is defined as the expected area of  $\mathbf{T}^{(2)}(\Delta d)$  in a unit square:

$$\mu_2(\mathbf{T}^{(2)}(\Delta d)) = \mathbb{E} \left( \left| \mathbf{T}^{(2)}(\Delta d) \cap [0, 1]^2 \right| \right), \quad (9)$$

where  $|S|$  denotes the area of  $S$  (i.e., two-dimensional Lebesgue measure of  $S$ ). Similarly, the area intensity of  $\mathbf{T}_{kj}^{(2)}(\Delta d)$  is

$$\mu_2(\mathbf{T}_{kj}^{(2)}(\Delta d)) = \mathbb{E} \left( \left| \mathbf{T}_{kj}^{(2)}(\Delta d) \cap [0, 1]^2 \right| \right). \quad (10)$$

<sup>1</sup>Because  $\Phi_1, \dots, \Phi_K$  are stationary,  $\mathbf{T}^{(1)}$  is also stationary, and thus the unit square could be arbitrarily picked on  $\mathbb{R}^2$ .

Because  $\Phi_1, \Phi_2, \dots, \Phi_K$  are stationary and isotropic,  $\mathbf{T}^{(2)}(\Delta d)$  and  $\mathbf{T}_{kj}^{(2)}(\Delta d)$  are also stationary and isotropic. As a result, given a reference UE located at  $\mathbf{0}$ , the area intensity of  $\mathbf{T}^{(2)}(\Delta d)$  (resp.  $\mathbf{T}_{kj}^{(2)}(\Delta d)$ ) is equal to the probability that the reference UE at  $\mathbf{0}$  is in  $\mathbf{T}^{(2)}(\Delta d)$  (resp.  $\mathbf{T}_{kj}^{(2)}(\Delta d)$ ).

$$\mu_2(\mathbf{T}^{(2)}(\Delta d)) = \mathbb{P}(\mathbf{0} \in \mathbf{T}^{(2)}(\Delta d)), \quad (11)$$

$$\mu_2(\mathbf{T}_{kj}^{(2)}(\Delta d)) = \mathbb{P}(\mathbf{0} \in \mathbf{T}_{kj}^{(2)}(\Delta d)). \quad (12)$$

We observe that the probabilities in (11) and (12) are analytically tractable, which will be presented in the next subsection.

## 4.3 Derivations of the Area Intensities

In this subsection, we present the derivations of  $\mathbb{P}(\mathbf{0} \in \mathbf{T}^{(2)}(\Delta d))$  and  $\mathbb{P}(\mathbf{0} \in \mathbf{T}_{kj}^{(2)}(\Delta d))$ . First, we study the probability that the reference UE at  $\mathbf{0}$  is in  $\mathbf{T}_{kj}^{(2)}(\Delta d)$ , given that it is associated to a tier- $k$  BS at a distance of  $r_0$  from it. By employing both analytic geometric and stochastic geometric tools, we derive the following theorem:

**THEOREM 1.** *Suppose the reference UE is located at  $\mathbf{0}$ , it is associated with a tier- $k$  BS, and their distance is  $R$ . The conditional probability that  $\mathbf{0} \in \mathbf{T}_{kj}^{(2)}(\Delta d)$  given  $R = r_0$  is*

$$\mathbb{P} \left( \mathbf{0} \in \mathbf{T}_{kj}^{(2)}(\Delta d) \mid R = r_0, \text{ tier} = k \right) = 1 - \exp \left( -2\lambda_j \Delta d r_0 \mathcal{F}(\beta_{kj}) + \mathcal{O}(\Delta d^2) \right), \quad (13)$$

where

$$\mathcal{F}(\beta) \triangleq \frac{1}{\beta^2} \int_0^\pi \sqrt{(\beta^2 + 1) - 2\beta \cos(\theta)} d\theta. \quad (14)$$

The proof is omitted due to the limited space.

Second, through stochastic geometric tools and deconditioning on  $R$ , we can derive the unconditioned probabilities that the reference UE at  $\mathbf{0}$  is in  $\mathbf{T}^{(2)}(\Delta d)$  and in  $\mathbf{T}_{kj}^{(2)}(\Delta d)$ :

**THEOREM 2.** *The area intensities of  $\mathbf{T}^{(2)}(\Delta d)$  and  $\mathbf{T}_{kj}^{(2)}(\Delta d)$  are:*

$$(a) \quad \mu_2(\mathbf{T}^{(2)}(\Delta d)) = \mathbb{P}(\mathbf{0} \in \mathbf{T}^{(2)}(\Delta d)) = \sum_{k=1}^K \frac{\lambda_k \left( \sum_{i=1}^K \lambda_i \Delta d \mathcal{F}(\beta_{ki}) \right)}{\left( \sum_{i=1}^K \lambda_i \beta_{ik}^2 \right)^{\frac{3}{2}}} + \mathcal{O}(\Delta d^2). \quad (15)$$

$$(b) \quad \mu_2(\mathbf{T}_{kj}^{(2)}(\Delta d)) = \mathbb{P}(\mathbf{0} \in \mathbf{T}_{kj}^{(2)}(\Delta d)) = \begin{cases} \frac{\lambda_k (\lambda_j \Delta d \mathcal{F}(\beta_{kj}))}{\left( \sum_{i=1}^K \lambda_i \beta_{ik}^2 \right)^{\frac{3}{2}}} + \frac{\lambda_j (\lambda_k \Delta d \mathcal{F}(\beta_{jk}))}{\left( \sum_{i=1}^K \lambda_i \beta_{ij}^2 \right)^{\frac{3}{2}}} + \mathcal{O}(\Delta d^2) & \text{if } k \neq j, \\ \frac{\lambda_k^2 \Delta d \mathcal{F}(1)}{\left( \sum_{i=1}^K \lambda_i \beta_{ik}^2 \right)^{\frac{3}{2}}} + \mathcal{O}(\Delta d^2) & \text{if } k = j. \end{cases} \quad (16)$$

See Appendix for the proof.

## 4.4 From Area Intensities to Handoff Rates

In this subsection, we derive handoff rates from area intensities derived in Theorem 2. This involves two steps:

(1) from area intensities  $\mu_2(\mathbf{T}^{(2)}(\Delta d))$  and  $\mu_2(\mathbf{T}_{kj}^{(2)}(\Delta d))$  to length intensities  $\mu_1(\mathbf{T}^{(1)})$  and  $\mu_1(\mathbf{T}_{kj}^{(1)})$ , and (2) from length intensities to handoff rates.

First, we derive the length intensity  $\mu_1(\mathbf{T}^{(1)})$  (resp.  $\mu_1(\mathbf{T}_{kj}^{(1)})$ ) from the area intensity  $\mu_2(\mathbf{T}^{(2)}(\Delta d))$  (resp.  $\mu_2(\mathbf{T}_{kj}^{(2)}(\Delta d))$ ) as follows

**THEOREM 3.** *The length intensities of  $\mathbf{T}^{(1)}$  and  $\mathbf{T}_{kj}^{(1)}$  can be computed as follows:*

(a)

$$\mu_1(\mathbf{T}^{(1)}) = \sum_{k=1}^K \frac{\lambda_k \left( \sum_{i=1}^K \lambda_i \mathcal{F}(\beta_{ki}) \right)}{2 \left( \sum_{i=1}^K \lambda_i \beta_{ik}^2 \right)^{\frac{3}{2}}}. \quad (17)$$

(b)

$$\mu_1(\mathbf{T}_{kj}^{(1)}) = \begin{cases} \frac{\lambda_k \lambda_j \mathcal{F}(\beta_{kj})}{2 \left( \sum_{i=1}^K \lambda_i \beta_{ik}^2 \right)^{\frac{3}{2}}} + \frac{\lambda_j \lambda_k \mathcal{F}(\beta_{jk})}{2 \left( \sum_{i=1}^K \lambda_i \beta_{ij}^2 \right)^{\frac{3}{2}}} & \text{if } k \neq j, \\ \frac{\lambda_k^2 \mathcal{F}(1)}{2 \left( \sum_{i=1}^K \lambda_i \beta_{ik}^2 \right)^{\frac{3}{2}}} & \text{if } k = j. \end{cases} \quad (18)$$

**PROOF.** It follows Section 3.2 in [20] and [13] by taking  $\Delta d \rightarrow 0$ .  $\square$

**REMARK 1.** *Note that, if we consider the single-tier case by taking  $K = 1$ , we have  $\mathcal{F}(1) = 4$ , and  $\mu_1(\mathbf{T}^{(1)}) = \mu_1(\mathbf{T}_{11}^{(1)}) = 2\sqrt{\lambda_1}$ . This matches the length intensity of a standard Poisson Voronoi. See Section 10.6 of [43].*

Second, we can derive the expected number of handoffs of an active UE as follows:

**THEOREM 4.** *Let  $\mathcal{T}_0$  denote an arbitrary UE's trajectory on  $\mathbb{R}^2$  with length  $|\mathcal{T}_0|_1$ . Then, the expected number of intersections of  $\mathcal{T}_0$  and  $\mathbf{T}^{(1)}$  (resp.  $\mathbf{T}_{kj}^{(1)}$ ) are*

$$\mathbb{E} \left( \mathcal{N}(\mathcal{T}_0, \mathbf{T}^{(1)}) \right) = \frac{2}{\pi} \mu_1(\mathbf{T}^{(1)}) |\mathcal{T}_0|_1, \quad (19)$$

$$\mathbb{E} \left( \mathcal{N}(\mathcal{T}_0, \mathbf{T}_{kj}^{(1)}) \right) = \frac{2}{\pi} \mu_1(\mathbf{T}_{kj}^{(1)}) |\mathcal{T}_0|_1, \quad (20)$$

and the expected number of type  $k$ - $j$  handoffs are

$$\mathbb{E} \left( \mathcal{N}_{kj}(\mathcal{T}_0, \mathbf{T}_{kj}^{(1)}) \right) = \begin{cases} \frac{1}{2} \mathbb{E} \left( \mathcal{N}(\mathcal{T}_0, \mathbf{T}_{kj}^{(1)}) \right) & \text{if } k \neq j, \\ \mathbb{E} \left( \mathcal{N}(\mathcal{T}_0, \mathbf{T}_{kj}^{(1)}) \right) & \text{if } k = j. \end{cases} \quad (21)$$

**PROOF.**  $\mathbf{T}^{(1)}$  and  $\mathbf{T}_{kj}^{(1)}$  are stationary and isotropic fibre processes with length intensity  $\mu_1(\mathbf{T}^{(1)})$  and  $\mu_1(\mathbf{T}_{kj}^{(1)})$  respectively. The proof follows the conclusions in Section 9.3 of [43].  $\square$

Note that the expected number of type  $k$ - $j$  handoffs is the same as the expected number of type  $j$ - $k$  handoffs, both of which are equal to half of  $\mathbb{E} \left( \mathcal{N}(\mathcal{T}_0, \mathbf{T}_{kj}^{(1)}) \right)$ .

Let  $v$  denote the instantaneous velocity of an active UE,  $H(v)$  denote its overall handoff rate (i.e., sum handoff rate of all types), and  $H_{kj}(v)$  denote its type  $k$ - $j$  handoff rate. Then we have the following Corollary from Theorem 4:

**COROLLARY 1.**

$$H(v) = \frac{2}{\pi} \mu_1(\mathbf{T}^{(1)}) v, \quad (22)$$

$$H_{kj}(v) = \begin{cases} \frac{1}{\pi} \mu_1(\mathbf{T}_{kj}^{(1)}) v & \text{if } k \neq j, \\ \frac{2}{\pi} \mu_1(\mathbf{T}_{kj}^{(1)}) v & \text{if } k = j. \end{cases} \quad (23)$$

Note that the above handoff rates are instantaneous rates. Our analysis allows time-varying velocity for the UEs, in which case the handoff rates are also time varying.

## 5. EXPERIMENTAL STUDY

In this section, our analysis is validated via experimenting with real-world traces and simulations.

### 5.1 Yonsei Trace Data

We use the real-world Yonsei Trace [17] to validate our analytical results. The trace was accumulated from 12 commercial mobile phones during an 8-month period in 2011 in the city of Seoul. An application named SmartDC had been running on the commercial mobile phones equipped with GPS, GSM, and WiFi. For every 2 to 5 minutes, the application collected UE's location information (latitude and longitude), the MAC addresses of surrounding WiFi APs, and the cell IDs of nearby cellular BSs they could detect. Each AP has a unique MAC address and each BS has a unique cell ID. By analyzing the data set, we are able to determine which APs and BSs a UE could detect at the recorded coordinates and time instants. In the following, we regard cellular BSs as tier-1 BSs and APs as tier-2 BSs.

### 5.2 Data Processing

#### 5.2.1 Location Approximations of APs and BSs

As the data set does not explicitly provide the latitudes and longitudes of APs and BSs, we apply the following approach to approximate their locations: for each AP (resp. BS), we list all the coordinates recorded by UEs when they are able to detect the AP (resp. BS). Then, we approximate the location of the AP (resp. BS), by taking the average of these recorded coordinates.

#### 5.2.2 Reference Region

In order to avoid the edge effect, we define a reference region, in which most recorded coordinates are located. The UEs' trajectories are only accounted inside the reference region. By plotting the cumulative distribution function (cdf) of the latitude (resp. longitude) of all recorded coordinates, we observe a sharp step upward between  $37.48^\circ N$  and  $37.58^\circ N$  (resp.  $126.9^\circ E$  and  $127.1^\circ E$ ). As a consequence, we employ the rectangle defined by  $37.48^\circ N$ , and  $37.58^\circ N$ ,  $126.9^\circ E$ , and  $127.1^\circ E$  as the reference region.

#### 5.2.3 UE Trajectory

In the trace data, the coordinates of a UE are recorded only once every few minutes. To recover its full trajectory, we regard it as moving in a straight line at a constant velocity between two consecutive recorded coordinates. Thus, interpolations can be made to determine the coordinate of the UE at any time. Note that only the trajectories inside the reference region are used.

#### 5.2.4 Handoff Rates

Through the locations of BSs and APs, as well as the UE trajectories, we are able to derive all types of empirical handoff rates following the biased user association scheme discussed in Section 3.2. If we ignore all the APs, we can also derive the empirical handoff rates for one-tier case.

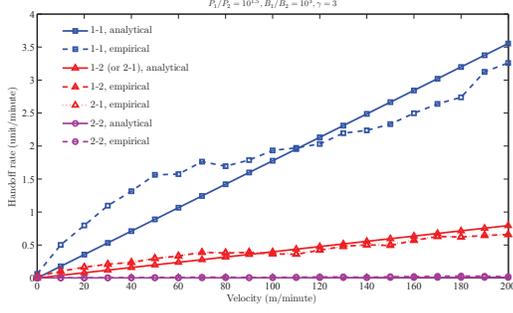


Figure 3: Two-tier case: comparison of analytical and empirical handoff rates.

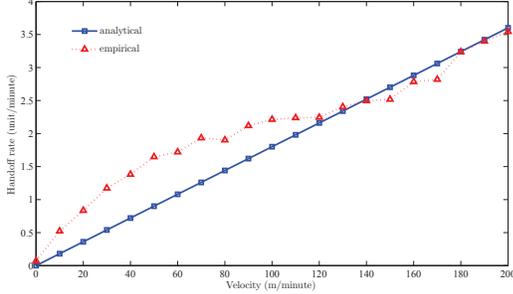


Figure 4: One-tier case: comparison of analytical and empirical handoff rates.

### 5.2.5 BS and AP Intensities

The AP (resp. BS) density is computed as the number of APs (resp. BSs) over the area of the reference region, which is 455.1 unit/km<sup>2</sup> (resp. 52.6 unit/km<sup>2</sup>). This indicates an urban area with high population and BS densities.

## 5.3 Empirical Results

We compare the handoff rates derived from our analysis and those from our empirical study based on the Yonsei Trace. The empirical handoff rates are derived from the steps in Sections 5.2.1 - 5.2.4. For the analytical results, we use the BS and AP intensities shown in Section 5.2.5 as input parameters.

For the two-tier case, the comparison of analytical and empirical handoff rates is shown in Fig. 3. For the one-tier case (by eliminating all the APs), the comparison is shown in Fig. 4. Both figures illustrate the accuracy of our analysis. When the UE's velocity is low, empirical handoff rates are slightly greater than analytical handoff rates. This is because the locations of APs and BSs are not strictly homogeneous distributed (e.g., some APs and BSs are crowded along some streets, or at the center of the urban region). We also observe that UEs with lower velocity are more likely to be sampled in the region with higher AP and BS densities. As a consequence, the empirical handoff rates are higher than those expected by our analytical results.

Fig. 3 and Fig. 4 also show that type 1-1 horizontal handoff rates are almost the same in the one-tier and two-tier cases, but extra type 1-2 and type 2-1 vertical handoffs are introduced in the two-tier case. This agrees with our expectation that adding a second tier of APs brings more vertical handoffs. In addition, as a validation of (21), type 1-2 and type 2-1 handoff rates are almost the same in empirical results.

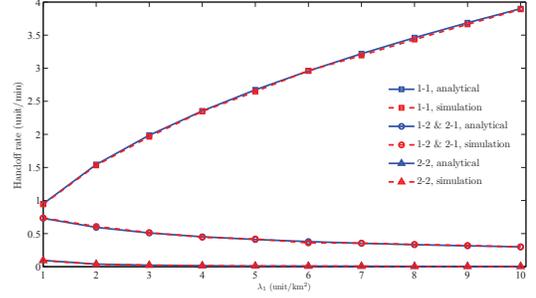


Figure 5: Two-tier case: handoff rates under different  $\lambda_1$ .

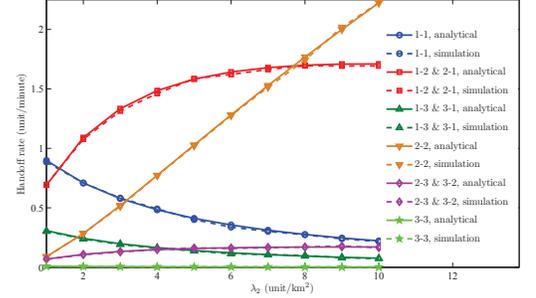


Figure 6: Three-tier case: handoff rates under different  $\lambda_2$ .

## 5.4 Simulation Study

In this subsection, we present simulation results to further demonstrate our analysis in more complex heterogeneous cellular networks.

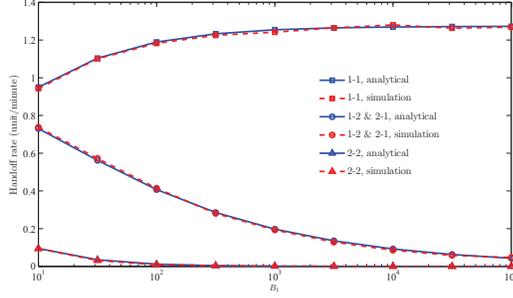
### 5.4.1 Simulation Setup

The simulation procedure is as follows: in each round of simulation, two or three tiers of BSs are generated on a 10 km  $\times$  10 km square. Then, we randomly generate 5 waypoints  $X_1, \dots, X_5$  in the central 5 km  $\times$  5 km square (uniformly distributed). The five line segments  $X_1X_2, X_2X_3, \dots, X_4X_5$  construct the trajectory of an active UE. In this way, we derive the simulated handoff rates in this round of simulation. The above procedure is repeated 200 rounds to derive one simulated data point. Note that in this subsection, in order to avoid overlapping in figures, we only show the sum rate of type  $j-k$  and type  $k-j$  ( $k \neq j$ ) handoffs for easier inspection; the individual handoff rates are half of the sum handoff rate.

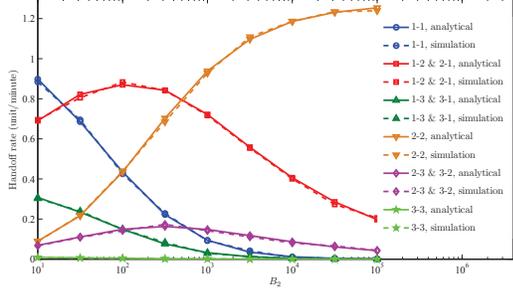
### 5.4.2 Handoff Rates under Different BS Intensities

We study handoff rates under different BS intensities. Fig. 5 shows a two-tier case, with parameters as follows:  $P_1 = 30$  dBm,  $P_2 = 20$  dBm, and  $B_1 = B_2 = 1$ ,  $\lambda_2 = 1$  unit/km<sup>2</sup>. Fig. 6 shows a three-tier case, with parameters as follows:  $P_1 = 30$  dBm,  $P_2 = 20$  dBm,  $P_3 = 10$  dBm,  $B_1 = B_2 = B_3 = 1$ , and  $\lambda_1 = \lambda_3 = 1$  unit/km<sup>2</sup>. The parameter values  $\gamma = 3$  and  $v = 60$  km/h are used for both Fig. 5 and Fig. 6.

Fig. 5 illustrates that increasing  $\lambda_1$  leads to higher type 1-1 handoff rate but lower type 2-2 handoff rate. Fig. 6 illustrates that increasing  $\lambda_2$  leads to higher type 2-2 handoff rate but lower type 1-1 and 1-3 & 3-1 handoff rates. Both observations suggest that increasing the BS intensity



**Figure 7: Two-tier case: handoff rates under different  $B_1$ .**



**Figure 8: Three-tier case: handoff rates under different  $B_2$ .**

of one tier causes higher horizontal handoff rate within this tier, but lower handoff rates outside this tier.

### 5.4.3 Handoff Rates under Different Association Bias Values

Next, we study handoff rates under different association bias values. Fig. 7 shows a two-tier case, with parameters as follows:  $P_1 = 30$  dBm,  $P_2 = 20$  dBm,  $B_2 = 1$ , and  $\lambda_1 = \lambda_2 = 1$  unit/km<sup>2</sup>. Fig. 8 shows a three-tier case, with parameters as follows:  $P_1 = 30$  dBm,  $P_2 = 20$  dBm,  $P_3 = 10$  dBm,  $B_1 = B_3 = 1$ ,  $\lambda_1 = \lambda_2 = \lambda_3 = 1$  unit/km<sup>2</sup>. The parameter values  $\gamma = 3$  and  $v = 60$  km/h are used for both Fig. 7 and Fig. 8. These figures suggest that, increasing the association bias value of one tier has a similar effect as increasing the BS intensity of this tier, leading to higher horizontal handoff rate within this tier, but lower handoff rates outside this tier.

## 6. CONCLUSIONS

In this work, we provide a theoretical framework to study user mobility in heterogeneous multi-tier cellular networks. Through establishing a stochastic geometric framework, we fully capture the irregularly shaped network topologies introduced by the small-cell structure. Theoretical expressions for the rates of all types of handoffs experienced by an active UE with arbitrary movement trajectory are derived. Empirical study on the Yonsei Trace and extensive simulation are conducted, validating the accuracy and usefulness of our analytical conclusions.

## 7. ACKNOWLEDGEMENTS

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## 8. APPENDIX

Proof. (a) Let  $E_i$  denote the event that there is at least one tier- $i$  BS located in  $S_{ki}(\Delta d)$ . Then

$$\begin{aligned}
 & \mathbb{P}(\mathbf{0} \in \mathbf{T}^{(2)}(\Delta d) | R = r_0, \text{tier} = k) \\
 &= 1 - \mathbb{P}(\overline{E_1} \cap \overline{E_2} \cap \dots \cap \overline{E_K} | R = r_0, \text{tier} = k) \\
 &= 1 - \exp\left(-\sum_{i=1}^K |S_{ki}(\Delta d)| \lambda_i\right) \\
 &= 1 - \exp\left(-\sum_{i=1}^K 2\lambda_i \Delta d r_0 \mathcal{F}(\beta_{ki}) + \mathcal{O}(\Delta d^2)\right) \\
 &= \sum_{i=1}^K 2\lambda_i \Delta d r_0 \mathcal{F}(\beta_{ki}) + \mathcal{O}(\Delta d^2). \tag{24}
 \end{aligned}$$

Furthermore, according to the results in [27], the probability density function (pdf) of the distance between the reference UE and the reference BS is

$$f_k(r_0 | \text{tier} = k) = \frac{2\pi\lambda_k}{A_k} r_0 \exp\left(-\pi r_0^2 \sum_{i=1}^K \lambda_i \beta_{ik}^2\right). \tag{25}$$

Also, we have  $\mathbb{P}(\text{tier} = k) = A_k$ , thus

$$\begin{aligned}
 & \mathbb{P}(\mathbf{0} \in \mathbf{T}^{(2)}(\Delta d)) \\
 &= \sum_{k=1}^K \int_0^\infty \mathbb{P}(\mathbf{0} \in \mathbf{T}^{(2)}(\Delta d) | R = r_0, \text{tier} = k) \\
 & \quad \cdot f_k(r_0 | \text{tier} = k) \mathbb{P}(\text{tier} = k) dr_0 \\
 &= \sum_{k=1}^K \int_0^\infty 2\pi\lambda_k r_0 \exp\left(-\pi r_0^2 \sum_{i=1}^K \lambda_i \beta_{ik}^2\right) \\
 & \quad \cdot \left(\sum_{i=1}^K 2\lambda_i \Delta d r_0 \mathcal{F}(\beta_{ki}) + \mathcal{O}(\Delta d^2)\right) dr_0 \\
 &= \sum_{k=1}^K \frac{\lambda_k \left(\sum_{i=1}^K \lambda_i \Delta d \mathcal{F}(\beta_{ki}) + \mathcal{O}(\Delta d^2)\right)}{\left(\sum_{i=1}^K \lambda_i \beta_{ik}^2\right)^{\frac{3}{2}}}, \tag{26}
 \end{aligned}$$

which completes the proof of (a).

(b)

Similar to (26), if  $k \neq j$  we have

$$\begin{aligned}
 & \mathbb{P}(\mathbf{0} \in \mathbf{T}_{kj}^{(2)}(\Delta d)) \\
 &= \int_0^\infty \mathbb{P}(\mathbf{0} \in \mathbf{T}_{kj}^{(2)}(\Delta d) | R = r_0, \text{tier} = k) f_k(r_0 | \text{tier} = k) \\
 & \quad \cdot \mathbb{P}(\text{tier} = k) dr_0 \\
 & \quad + \int_0^\infty \mathbb{P}(\mathbf{0} \in \mathbf{T}_{kj}^{(2)}(\Delta d) | R = r_0, \text{tier} = j) f_j(r_0 | \text{tier} = j) \\
 & \quad \cdot \mathbb{P}(\text{tier} = j) dr_0 \\
 &= \frac{\lambda_k (\lambda_j \Delta d \mathcal{F}(\beta_{kj}) + \mathcal{O}(\Delta d^2))}{\left(\sum_{i=1}^K \lambda_i \beta_{ik}^2\right)^{\frac{3}{2}}} + \frac{\lambda_j (\lambda_k \Delta d \mathcal{F}(\beta_{jk}) + \mathcal{O}(\Delta d^2))}{\left(\sum_{i=1}^K \lambda_i \beta_{ij}^2\right)^{\frac{3}{2}}}. \tag{27}
 \end{aligned}$$

Otherwise, if  $k = j$  we have

$$\mathbb{P}(\mathbf{0} \in \mathbf{T}_{kk}^{(2)}(\Delta d)) = \frac{\lambda_k (\lambda_k \Delta d \mathcal{F}(1) + \mathcal{O}(\Delta d^2))}{\left(\sum_{i=1}^K \lambda_i \beta_{ik}^2\right)^{\frac{3}{2}}}, \tag{28}$$

which completes the proof of (b).  $\square$

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